

$$\begin{aligned}
\prod_{\lambda=1}^{N/2} g_{\lambda} = & \exp \left\{ -\frac{1}{2}(\sigma_1^2 + \sigma_2^2 + \sigma_3^2 + \sigma_4^2 + \sigma_{12}^2 + \sigma_{23}^2 + \sigma_{31}^2) \right. \\
& - \frac{i}{N^{1/2}} (\sigma_1 \sigma_4 \sigma_{23} + \sigma_1 \sigma_3 \sigma_{31} + \sigma_1 \sigma_2 \sigma_{12} + \sigma_2 \sigma_3 \sigma_{23} \\
& + \sigma_2 \sigma_4 \sigma_{31} + \sigma_3 \sigma_4 \sigma_{12}) \\
& + \frac{1}{N} (\sigma_1 \sigma_2 \sigma_3 \sigma_4 + \sigma_2 \sigma_3 \sigma_{31} \sigma_{12} + \sigma_2 \sigma_4 \sigma_{12} \sigma_{23} + \sigma_3 \sigma_4 \sigma_{23} \sigma_{31} \\
& \left. + \sigma_1 \sigma_4 \sigma_{31} \sigma_{12} + \sigma_1 \sigma_3 \sigma_{12} \sigma_{23} + \sigma_1 \sigma_2 \sigma_{23} \sigma_{31}) \right\} \\
& \times \left\{ 1 + O\left(\frac{1}{N}\right) \right\} \quad (\text{II.1})
\end{aligned}$$

where $O(1/N)$ represents the terms of order $1/N$ or higher in which the terms of order $1/N$ contain only even powers of the σ 's.

APPENDIX III

Evaluating the sevenfold integral (2.3)

III. 1. The σ_1 integration

Substitute for $\prod_{\lambda=1}^{N/2} g_{\lambda}$ from (II.1) into (2.3), combine the terms in the exponent involving σ_1 , complete the square, and perform the σ_1 integration.

III. 2. The remaining integrations

One continues in this way, carrying out the successive integrations with respect to $\sigma_2, \sigma_3, \dots$, until finally (2.5) is obtained.

It is instructive to compare these integrations with those of the earlier paper (Appendix IV).

References

- DEITTA, G., EDMONDS, J., LANGS, D. & HAUPTMAN, H. (1974). Amer. Cryst. Assoc. Summer Meeting at Penn State Univ. Abstract D3.
- DEITTA, G., EDMONDS, J., LANGS, D. & HAUPTMAN, H. (1975). *Acta Cryst.* A31, 472–479.
- EINSPAHR, H., GARTLAND, G., FREEMAN, G. & SCHENCK, H. (1974). Amer. Cryst. Assoc. Summer Meeting at Penn State Univ. Abstract B6.
- HAUPTMAN, H. (1975a). *Acta Cryst.* A31, 671–679.
- HAUPTMAN, H. (1975b). *Acta Cryst.* A31, 680–687.
- HAUPTMAN, H., FISHER, J., HANCOCK, H. & NORTON, D. (1969). *Acta Cryst.* B25, 811–814.
- HAUPTMAN, H. & GREEN, E. A. (1976). *Acta Cryst.* A32, 45–49.
- KARLE, J. J. & HAUPTMAN, H. (1958). *Acta Cryst.* 11, 264–269.

Acta Cryst. (1976). A32, 45

Conditional Probability Distributions of the Four-Phase Structure Invariant

$\varphi_{\mathbf{h}} + \varphi_{\mathbf{k}} + \varphi_{\mathbf{l}} + \varphi_{\mathbf{m}}$ in $P\bar{1}^*$

BY HERBERT HAUPTMAN AND EDWARD A. GREEN

Medical Foundation of Buffalo, 73 High Street, Buffalo, New York 14203, U.S.A.

(Received 16 June 1975; accepted 28 July 1975)

A crystal structure in $P\bar{1}$ is assumed to be fixed and the seven non-negative numbers $R_1, R_2, R_3, R_4, R_{12}, R_{23}, R_{31}$ are also given. It is assumed that $\mathbf{h}, \mathbf{k}, \mathbf{l}, \mathbf{m}$ are random variables uniformly and independently distributed over the subsets of reciprocal space defined by

$$|E_{\mathbf{h}}| = R_1, |E_{\mathbf{k}}| = R_2, |E_{\mathbf{l}}| = R_3, |E_{\mathbf{m}}| = R_4, \quad (1)$$

$$|E_{\mathbf{h}+\mathbf{k}}| = R_{12}, |E_{\mathbf{k}+\mathbf{l}}| = R_{23}, |E_{\mathbf{l}+\mathbf{h}}| = R_{31}, \quad (2)$$

and

$$\mathbf{h} + \mathbf{k} + \mathbf{l} + \mathbf{m} = 0. \quad (3)$$

Then the structure invariant

$$\varphi = \varphi_{\mathbf{h}} + \varphi_{\mathbf{k}} + \varphi_{\mathbf{l}} + \varphi_{\mathbf{m}} \quad (4)$$

is a function of the primitive random variables $\mathbf{h}, \mathbf{k}, \mathbf{l}, \mathbf{m}$. The conditional probability distribution of φ , given (1) and (2), is obtained and compared with the conditional probability distribution of φ when only (1) is given. Some calculations are presented which show the usefulness of the distribution, given (1) and (2), in estimating the value of φ .

1. Introduction

The methods introduced in two previous papers (Hauptman 1975a, b) for $P1$ are applied here to the space

group $P\bar{1}$. Again, as in the earlier work, the joint probability distribution of seven structure factors [Green & Hauptman, 1976, equation (2.5)] leads directly to the conditional probability distribution of the four-phase structure invariant $\varphi_{\mathbf{h}} + \varphi_{\mathbf{k}} + \varphi_{\mathbf{l}} + \varphi_{\mathbf{m}}$, given the seven magnitudes $|E_{\mathbf{h}}|, |E_{\mathbf{k}}|, |E_{\mathbf{l}}|, |E_{\mathbf{m}}|, |E_{\mathbf{h}+\mathbf{k}}|, |E_{\mathbf{k}+\mathbf{l}}|, |E_{\mathbf{l}+\mathbf{h}}|$. However, in contrast to the earlier distri-

* Presented at the Charlottesville meeting of the ACA, March 9–13, 1975, Abstract A1.

bution which is continuous, the distribution to be derived here is discrete because the structure invariant has only the two possible values, 0 or π .

2. The joint conditional probability distribution of the four phases $\varphi_{\mathbf{h}}$, $\varphi_{\mathbf{k}}$, $\varphi_{\mathbf{l}}$, $\varphi_{\mathbf{m}}$, given the seven magnitudes, $|E_{\mathbf{h}}|$, $|E_{\mathbf{k}}|$, $|E_{\mathbf{l}}|$, $|E_{\mathbf{m}}|$, $|E_{\mathbf{h}+\mathbf{k}}|$, $|E_{\mathbf{k}+\mathbf{l}}|$, $|E_{\mathbf{l}+\mathbf{h}}|$

Suppose that a crystal structure consisting of N identical atoms per unit cell in the space group $P\bar{1}$ is specified and that the seven non-negative numbers $R_1, R_2, R_3, R_4, R_{12}, R_{23}, R_{31}$ are also fixed. Denote by $\varphi_{\mathbf{h}}$ the phase of the normalized structure factor $E_{\mathbf{h}}$. Define the four-fold Cartesian product $W \times W \times W \times W$ of reciprocal space W with itself to consist of the collection of all ordered quadruples $(\mathbf{h}, \mathbf{k}, \mathbf{l}, \mathbf{m})$ of reciprocal vectors $\mathbf{h}, \mathbf{k}, \mathbf{l}, \mathbf{m}$. Suppose finally that the ordered quadruple $(\mathbf{h}, \mathbf{k}, \mathbf{l}, \mathbf{m})$ of reciprocal vectors is a random variable which is uniformly distributed over the subset of the Cartesian product $W \times W \times W \times W$ for which

$$|E_{\mathbf{h}}| = R_1, |E_{\mathbf{k}}| = R_2, |E_{\mathbf{l}}| = R_3, |E_{\mathbf{m}}| = R_4, \quad (2.1)$$

$$|E_{\mathbf{h}+\mathbf{k}}| = R_{12}, |E_{\mathbf{k}+\mathbf{l}}| = R_{23}, |E_{\mathbf{l}+\mathbf{h}}| = R_{31}, \quad (2.2)$$

and

$$\mathbf{h} + \mathbf{k} + \mathbf{l} + \mathbf{m} = 0. \quad (2.3)$$

[Strictly speaking, in order to insure that the range of the primitive random variable $(\mathbf{h}, \mathbf{k}, \mathbf{l}, \mathbf{m})$ be non-vacuous, it is necessary, for example, to replace the equality $|E_{\mathbf{h}}| = R_1$ of (2.1) by the inequalities $R_1 \leq |E_{\mathbf{h}}| \leq R_1 + dR_1$, where dR_1 is a 'small' positive number, *etc.*]. In view of (2.1)–(2.3), the random variables $\mathbf{h}, \mathbf{k}, \mathbf{l}, \mathbf{m}$, the components of the ordered quadruple $(\mathbf{h}, \mathbf{k}, \mathbf{l}, \mathbf{m})$, are not independently distributed in reciprocal space. Then $\varphi_{\mathbf{h}}, \varphi_{\mathbf{k}}, \varphi_{\mathbf{l}}, \varphi_{\mathbf{m}}$, the phases of the normalized structure factors $E_{\mathbf{h}}, E_{\mathbf{k}}, E_{\mathbf{l}}, E_{\mathbf{m}}$, as functions of the primitive random variables $\mathbf{h}, \mathbf{k}, \mathbf{l}, \mathbf{m}$, are themselves random variables. Denote by $P(\Phi_1, \Phi_2, \Phi_3, \Phi_4 | R_1, R_2, R_3, R_4, R_{12}, R_{23}, R_{31})$ the joint conditional probability distribution of the four phases $\varphi_{\mathbf{h}}, \varphi_{\mathbf{k}}, \varphi_{\mathbf{l}}, \varphi_{\mathbf{m}}$, given (2.1), (2.2) and (2.3). Then $P(\Phi_1, \Phi_2, \Phi_3, \Phi_4 | R_1, R_2, R_3, R_4, R_{12}, R_{23}, R_{31})$ is found from (2.5) of the previous paper (Green & Hauptman, 1976) by fixing the magnitudes of $S_1, S_2, S_3, S_4, S_{12}, S_{23}, S_{31}$ in accordance with the scheme

$$|S_1| = R_1, |S_2| = R_2, |S_3| = R_3, |S_4| = R_4, \quad (2.4)$$

$$|S_{12}| = R_{12}, |S_{23}| = R_{23}, |S_{31}| = R_{31}, \quad (2.5)$$

i.e.

$$S_1 = R_1 \cos \Phi_1, S_2 = R_2 \cos \Phi_2, \\ S_3 = R_3 \cos \Phi_3, S_4 = R_4 \cos \Phi_4, \quad (2.6)$$

$$S_{12} = R_{12} \cos \Phi_{12}, S_{23} = R_{23} \cos \Phi_{23}, \\ S_{31} = R_{31} \cos \Phi_{31}, \quad (2.7)$$

where Φ_{12} is the variable associated with the phase $\varphi_{\mathbf{h}+\mathbf{k}}$, *etc.*, summing with respect to S_{12}, S_{23}, S_{31} over their two possible signs (+ and -) or, equivalently, sum-

ming with respect to $\Phi_{12}, \Phi_{23}, \Phi_{31}$ over their two possible values (0 and π), and multiplying the result by a suitable normalizing factor. Carrying out these summations one finally obtains, correct up to and including terms of order $1/N$ [since $O(1/N)$ of the previous paper consists of all terms of order $1/N$ or higher in which the terms of order $1/N$ contain only even powers of the S 's],

$$P(\Phi_1, \Phi_2, \Phi_3, \Phi_4 | R_1, R_2, R_3, R_4, R_{12}, R_{23}, R_{31}) \\ \simeq \frac{1}{K} \exp \{ -B \cos (\Phi_1 + \Phi_2 + \Phi_3 + \Phi_4) \} \\ \times \cosh \frac{R_{12} Y_{12}}{N^{1/2}} \cosh \frac{R_{23} Y_{23}}{N^{1/2}} \cosh \frac{R_{31} Y_{31}}{N^{1/2}} \quad (2.8)$$

where

$$B = \frac{2}{N} R_1 R_2 R_3 R_4, \quad (2.9)$$

$$Y_{12} = [R_1^2 R_2^2 + R_3^2 R_4^2 + 2R_1 R_2 R_3 R_4 \\ \times \cos (\Phi_1 + \Phi_2 + \Phi_3 + \Phi_4)]^{1/2}, \quad (2.10)$$

$$Y_{23} = [R_2^2 R_3^2 + R_1^2 R_4^2 + 2R_1 R_2 R_3 R_4 \\ \times \cos (\Phi_1 + \Phi_2 + \Phi_3 + \Phi_4)]^{1/2}, \quad (2.11)$$

$$Y_{31} = [R_3^2 R_1^2 + R_2^2 R_4^2 + 2R_1 R_2 R_3 R_4 \\ \times \cos (\Phi_1 + \Phi_2 + \Phi_3 + \Phi_4)]^{1/2}, \quad (2.12)$$

and K is a suitable normalizing constant, independent of $\Phi_1, \Phi_2, \Phi_3, \Phi_4$. Although K is readily found by summing (2.8) over the 16 sets of values of $\Phi_1, \Phi_2, \Phi_3, \Phi_4$ and setting the result equal to unity, the value of this normalizing factor is not needed for the present purpose and is therefore not derived explicitly. Equations (2.8)–(2.12) should be compared with equations (2.5)–(2.9) of the earlier work in $P1$ (Hauptman, 1975*b*), but it should be emphasized again that the present distribution (2.8) is discrete since each of $\Phi_1, \Phi_2, \Phi_3, \Phi_4$ takes on only the two values 0, π .

It is clear from (2.8)–(2.12) that the distribution (2.8) is a function of the sum $\Phi = \Phi_1 + \Phi_2 + \Phi_3 + \Phi_4$. Hence (2.8) leads directly to the conditional distribution, given (2.1) and (2.2), of the sum $\varphi = \varphi_{\mathbf{h}} + \varphi_{\mathbf{k}} + \varphi_{\mathbf{l}} + \varphi_{\mathbf{m}}$, as is shown next.

3. The conditional probability distribution of the structure invariant $\varphi = \varphi_{\mathbf{h}} + \varphi_{\mathbf{k}} + \varphi_{\mathbf{l}} + \varphi_{\mathbf{m}}$, given the seven magnitudes $|E_{\mathbf{h}}|$, $|E_{\mathbf{k}}|$, $|E_{\mathbf{l}}|$, $|E_{\mathbf{m}}|$, $|E_{\mathbf{h}+\mathbf{k}}|$, $|E_{\mathbf{k}+\mathbf{l}}|$, $|E_{\mathbf{l}+\mathbf{h}}|$

Using the same hypotheses as in §2, the structure invariant

$$\varphi = \varphi_{\mathbf{h}} + \varphi_{\mathbf{k}} + \varphi_{\mathbf{l}} + \varphi_{\mathbf{m}} \quad (3.1)$$

is a random variable whose conditional probability distribution, given (2.1) and (2.2), $P(\Phi | R_1, R_2, R_3, R_4, R_{12}, R_{23}, R_{31})$, is readily found from (2.8)–(2.12). Thus, correct up to and including terms of order $1/N$, the major result of this paper is given by

$$\begin{aligned}
P(\Phi | R_1, R_2, R_3, R_4, R_{12}, R_{23}, R_{31}) \\
\simeq \frac{1}{L} \exp(-B \cos \Phi) \\
\times \cosh \frac{R_{12}Z_{12}}{N^{1/2}} \cosh \frac{R_{23}Z_{23}}{N^{1/2}} \cosh \frac{R_{31}Z_{31}}{N^{1/2}} \quad (3.2)
\end{aligned}$$

where B is defined by (2.9),

$$Z_{12} = [R_1^2 R_2^2 + R_3^2 R_4^2 + 2R_1 R_2 R_3 R_4 \cos \Phi]^{1/2}, \quad (3.3)$$

$$Z_{23} = [R_2^2 R_3^2 + R_1^2 R_4^2 + 2R_1 R_2 R_3 R_4 \cos \Phi]^{1/2}, \quad (3.4)$$

$$Z_{31} = [R_3^2 R_1^2 + R_2^2 R_4^2 + 2R_1 R_2 R_3 R_4 \cos \Phi]^{1/2}, \quad (3.5)$$

and the normalizing constant L is readily found to be

$$\begin{aligned}
L = \exp(-B) \cosh \frac{R_{12}Z_{12}^+}{N^{1/2}} \cosh \frac{R_{23}Z_{23}^+}{N^{1/2}} \cosh \frac{R_{31}Z_{31}^+}{N^{1/2}} \\
+ \exp(B) \cosh \frac{R_{12}Z_{12}^-}{N^{1/2}} \cosh \frac{R_{23}Z_{23}^-}{N^{1/2}} \cosh \frac{R_{31}Z_{31}^-}{N^{1/2}}, \quad (3.6)
\end{aligned}$$

where

$$Z_{12}^\pm = R_1 R_2 \pm R_3 R_4, \quad (3.7)$$

$$Z_{23}^\pm = R_2 R_3 \pm R_1 R_4, \quad (3.8)$$

$$Z_{31}^\pm = R_3 R_1 \pm R_2 R_4, \quad (3.9)$$

and the upper (lower) signs go together.

If one denotes by P_+ (P_-) the conditional probability, given (2.1) and (2.2), that

$$\varphi = \varphi_h + \varphi_k + \varphi_1 + \varphi_m = 0 \pmod{\pi}, \quad (3.10)$$

or that

$$\cos \varphi = +1 \quad (-1), \quad (3.11)$$

or that

$$E_h E_k E_1 E_m \text{ be positive (negative)}, \quad (3.12)$$

then (3.2) is replaced by the more suggestive

$$\begin{aligned}
P_\pm \simeq \frac{1}{L} \exp(\mp B) \cosh \frac{R_{12}Z_{12}^\pm}{N^{1/2}} \\
\times \cosh \frac{R_{23}Z_{23}^\pm}{N^{1/2}} \cosh \frac{R_{31}Z_{31}^\pm}{N^{1/2}}, \quad (3.13)
\end{aligned}$$

where the upper (lower) signs go together and L , Z_{12}^\pm , Z_{23}^\pm , Z_{31}^\pm are given by (3.6)–(3.9).

The similarity between (3.2) and (3.13) with the major result of the earlier work in $P1$ [equation (3.2), Hauptman, 1975b] is noteworthy. However, it must be stressed again that the present distributions are discrete, *i.e.* Φ must be 0 or π , while the earlier distribution is continuous.

It should be observed finally that P_+ may lie anywhere between 0 and 1; P_+ is close to 0 or close to 1 according as R_{12} , R_{23} , R_{31} are all relatively small or all relatively large respectively.

3.1. The conditional expected value and conditional variance of $\cos \varphi$

It is clear from (3.13) that the conditional expected value of $\cos \varphi$, given (2.1) and (2.2), is simply

$$\begin{aligned}
\epsilon = \epsilon(\cos \varphi) = \epsilon(\cos \varphi | R_1, R_2, R_3, R_4, R_{12}, R_{23}, R_{31}) \\
\simeq P_+ - P_- \quad (3.14)
\end{aligned}$$

Likewise the conditional variance of $\cos \varphi$, given (2.1) and (2.2), is found from (3.14) to be

$$\begin{aligned}
\sigma^2 = \sigma^2(\cos \varphi | R_1, R_2, R_3, R_4, R_{12}, R_{23}, R_{31}) \\
\simeq \epsilon(\cos^2 \varphi) - [\epsilon(\cos \varphi)]^2 = 4P_+ P_- \quad (3.15)
\end{aligned}$$

It follows from (3.15) that the conditional standard deviation is given by

$$\sigma = 2\sqrt{P_+ P_-} \quad (3.16)$$

3.2. The special case $|E_{h+k}| \simeq |E_{k+1}| \simeq |E_{1+h}| \simeq 0$

In the case that

$$|E_{h+k}| = R_{12} \simeq 0, \quad (3.17)$$

$$|E_{k+1}| = R_{23} \simeq 0, \quad (3.18)$$

$$|E_{1+h}| = R_{31} \simeq 0, \quad (3.19)$$

(3.13) is replaced by

$$P_\pm (|R_{12} \simeq R_{23} \simeq R_{31} \simeq 0) \simeq \frac{1}{L} \exp(\mp B) \quad (3.20)$$

where

$$L = 2 \cosh B \quad (3.21)$$

so that, in this special case, $P_+ < \frac{1}{2}$, φ is probably equal to π , and the larger the value of B the more likely it is that $\varphi = \pi$.

4. The conditional probability distribution of the structure invariant $\varphi = \varphi_h + \varphi_k + \varphi_1 + \varphi_m$, given the four magnitudes $|E_h|$, $|E_k|$, $|E_1|$, $|E_m|$

If, instead of being given the seven magnitudes (2.1) and (2.2), one is given only the four magnitudes (2.1), then the conditional probability distribution of φ , given (2.1), is found from (3.1) of the previous paper (Green & Hauptman, 1976) to be, correct up to and including terms of order $1/N$,

$$P_\pm(4) \simeq \frac{1}{M} \exp\left(\pm \frac{B}{2}\right), \quad (4.1)$$

where

$$M = 2 \cosh \frac{B}{2}, \quad (4.2)$$

and $P_+(4)[P_-(4)]$ is the conditional probability, given (2.1), that

$$\varphi = 0 \pmod{\pi} \quad (4.3)$$

or that

$$\cos \varphi = +1 \quad (-1). \quad (4.4)$$

It is noteworthy that, employing some 300–400 of the most reliably estimated values of the invariants listed in Table 1, unique values were obtained, with perfect accuracy, for 230 of the phases having $|E|$ values greater than 2. Thus this structure is solvable *via* estimated values of the four-phase structure invariants alone.

6. Concluding remarks

Conditional probability distributions of the four-phase structure invariant $\varphi = \varphi_{\mathbf{h}} + \varphi_{\mathbf{k}} + \varphi_{\mathbf{l}} + \varphi_{\mathbf{m}}$ in $P\bar{1}$, given, in the first instance only the four magnitudes (2.1) and, in the second instance, all seven magnitudes (2.1) and (2.2), have been found. The distributions lead to estimates for φ dependent on these magnitudes. The initial applications strongly suggest that the results secured here will find important application in the solution of complex crystal structures in $P\bar{1}$. It is suggested that the methods described here, themselves an extension of some recent work in $P1$, can be extended to treat structure invariants and seminvariants in general, and that the concept of ‘neighborhood of a structure invariant’, introduced in an earlier paper (Hauptman, 1975*b*) will play the central role in this development.

7. Comparison with a recent result of Giacovazzo (1975)

Just as this paper was being prepared for submission, a paper by Giacovazzo (1975) appeared covering material closely related to ours but employing a different probabilistic background and a different mathematical formalism. Hence a rare opportunity presented itself to compare the different approaches.

First, it should be emphasized that, in using the Klug formalism, Giacovazzo assumes implicitly that the reciprocal vectors $\mathbf{h}, \mathbf{k}, \mathbf{l}, \mathbf{m}$ are fixed and that atomic coordinates are the primitive random variables. This assumption contrasts sharply with the one made here which is that the crystal structure is fixed and the reciprocal vectors $\mathbf{h}, \mathbf{k}, \mathbf{l}, \mathbf{m}$ are the primitive random variables. Thus there are two kinds of probability distribution which are conceptually quite distinct (a point made many times previously, *e.g.* Hauptman, 1975*a*), and there is no reason to suppose that corresponding distributions will be identical. It therefore comes as no surprise that none of Giacovazzo’s distributions (11)–(13) agrees with our (3.13). In particular, Giacovazzo’s conclusion that ‘When N is large enough, (11) tells us that the product $E_1 E_2 E_3 E_7$ (our $E_{\mathbf{h}} E_{\mathbf{k}} E_{\mathbf{l}} E_{\mathbf{m}}$) is probably positive when $E_4^2 + E_5^2 + E_6^2 - 2 > 0$ (in our notation

$E_{\mathbf{h}+\mathbf{k}}^2 + E_{\mathbf{k}+\mathbf{l}}^2 + E_{\mathbf{l}+\mathbf{h}}^2 - 2 > 0$), probably negative when $E_4^2 + E_5^2 + E_6^2 < 2$.’ cannot possibly follow from our (3.13) which implies instead that whether $P_+ > \frac{1}{2}$ or $P_+ < \frac{1}{2}$ depends on whether an intricate interrelationship among all seven magnitudes (2.1) and (2.2) holds and not merely on a relationship among the three magnitudes (2.2) alone. Again, it should be stressed that it is our distribution which is the appropriate one for crystal-structure analysis [refer again to Hauptman (1975*a*)], since one is ordinarily given a fixed but unknown crystal structure, and structure-factor amplitudes are sampled from reciprocal space. Finally, we make no use of the Gram–Charlier expansion in our analysis, in contrast to Giacovazzo’s work, but employ instead a mathematical formalism only recently secured (*e.g.* Hauptman, 1975*a, b*). With this background then, the comparison between Giacovazzo’s result and ours is particularly illuminating.

Observe next that Giacovazzo’s chief result, (11) and (12), requires knowledge not only of seven magnitudes but of the sign of $E_4 E_5 E_6$ (our $E_{\mathbf{h}+\mathbf{k}} E_{\mathbf{k}+\mathbf{l}} E_{\mathbf{l}+\mathbf{h}}$) as well, whereas our (3.13) depends only on the seven magnitudes (2.1) and (2.2). Thus comparison is possible only between our (3.13) and Giacovazzo’s (13), the special case that

$$|E_{\mathbf{h}+\mathbf{k}} E_{\mathbf{k}+\mathbf{l}} E_{\mathbf{l}+\mathbf{h}}| \simeq 0, \quad (7.1)$$

since in this case the sign of $E_{\mathbf{h}+\mathbf{k}} E_{\mathbf{k}+\mathbf{l}} E_{\mathbf{l}+\mathbf{h}}$ is irrelevant.

Table 2 displays a large and representative sample of 40 invariants from Table 1 for which (7.1) is approximately satisfied, so that comparison between our (3.13) and Giacovazzo’s (13) is possible. The last two columns, headed $P(+)$ and $P(13)$, show the values of P_+ as calculated from our (3.13) and Giacovazzo’s (13) respectively. Comparison of the entries in these two columns reveals significant differences, and comparison with the true cosine values, $\cos(T)$, is particularly illuminating.

We wish to thank Dr. Charles Weeks who helped in the calculation of Table 1. The calculation of the values of 230 phases using Table 1 was carried out by Miss Mary Duffy to whom grateful acknowledgement is made. This research was supported by NIH Grant No. GM-19684 and NSF Grant No. MPS73-04992.

References

- GIACOVAZZO, C. (1975). *Acta Cryst.* **A31**, 252–259.
 GREEN, E. A. & HAUPTMAN, H. (1976). *Acta Cryst.* **A32**, 43–45.
 HAUPTMAN, H. (1975*a*). *Acta Cryst.* **A31**, 671–679.
 HAUPTMAN, H. (1975*b*). *Acta Cryst.* **A31**, 680–687.